

The Kumaraswamy-Rani Distribution and Its Applications

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Abstract

In this paper, a new distribution called the Kumaraswamy-Rani (KR) distribution, as a Special model from the class of Kumaraswamy Generalized (KW-G) distributions, is introduced. Its statistical properties are explored. Estimation parameters based on maximum likelihood are obtained. We have illustrated the performances of the proposed distribution by some simulation studies. Finally, the significance of the KR distribution in terms of modeling real data set has been highlighted before it has been compared to some one parameter lifetime distributions.

Keywords: Fixed point method • Kumaraswamy distribution • Maximum likelihood estimation • Newton-raphson method • Rani distribution

Introduction

Modeling and analysis of lifetime data is an important aspect of statistical work in such wide variety of scientific and technological fields as medicine, engineering, insurance and finance...

Therefore, the construction of the distributions is the most important work from a probabilistic point of view and recently many researchers are concentrated in this area [1-3].

More that, varying lifetime distributions and their generalizations have been created as exponential, Lindley, gamma, lognormal, Weibull, Sujatha, Akash, Rani. Further, A two-parameter family of distributions on called Kumaraswamy distribution is explored. It was introduced by Kumaraswamy [4]. It has the flexibility in modeling lifetime data.

The latter exhibits many similarities to the beta distribution and a number of advantages in terms of tractability which has been discussed by Jones [5]. Among these advantages, we mention: a simple explicit formula for its distribution function, its quantile function. L-moments and moments of order statistics [6]. In this direction, another new one parameter lifetime distribution has attracted the attention of statisticians and whetted then interest named Rani distribution. It was proposed by Rama Shanker and it has important statistical properties [7]. For the motivating properties of the Kumaraswamy and Rani distributions, we shall introduce

a new probability distribution based on the Kumaraswamy distribution and the Rani distribution. The proposed distribution is called the Kumaraswamy-Rani (KR) distribution with twos positive parameters denoted b and β . In this study some statistical properties of this distribution are derived to show their applicability in the real lifetime data analysis. The rest of the paper is organized as follows. In Section 2, the KR distribution is identified in addition, some important mathematical properties of the new distribution are presented. In Section 3, maximum likelihood estimation is discussed. In section 4, some simulation studies are reported. A real lifetime data set are used in Section 5 to illustrate the usefulness of the KR distribution. Finally, our work is crowned with conclusion.

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The KR Distribution

The Kumaraswamy-Generalized distribution

The cumulative density function (cdf) of the Kumaraswamy-Generalized (Kum-Generalized) distribution proposed by Cordeiro *et al.* [8] is given by

$$F(x) = 1 - (1 - (G(x))^a)^b, \quad (1)$$

Where $a > 0, b > 0$ are shape parameters and G is the cdf of a continuous random variable X . Consequently, the Kumaraswamy-Generalized distribution has the probability density function (pdf) specified by

$$f(x) = abg(x)G(x)^{a-1}(1 - G(x)^a)^{b-1}, \quad (2)$$

$$\text{where, } g(x) = \frac{\partial G(x)}{\partial x}.$$

According to Cordeiro *et al.* [8], using the binomial expansion for $b \in \mathbb{R}^+$, equation (2) can be rewritten as

$$f(x) = g(x) \sum_{i=0}^{\infty} w_i G(x)^{a(i+1)-1}, \quad (3)$$

where the binomial coefficient w_i is defined for all real numbers with,

$$w_i = (-1)^i ab \binom{b-1}{i}$$

The Rani distribution

The Rani distribution is proposed by Rama Shanker [7]. It is defined over the support $x > 0$ and it has the pdf given by:

$$f(x) = \frac{\theta^5}{\theta^5 + 24} (\theta + x^4) e^{-\theta x}, \quad (4)$$

where θ is the scale parameter of the Rani distribution denoted $R(\theta)$. The latter is a mixture of exponential (θ) and gamma ($5, \theta$) with mixing proportion $\frac{\theta^5}{\theta^5 + 24}$. Furthermore, the corresponding cumulative distribution function

(cdf) is given by:

$$F(x) = 1 - \left[1 + \frac{\theta x (\theta^3 x^3 + 4\theta^2 x^2 + 12\theta x + 24)}{\theta^5 + 24} \right] e^{-\theta x}. \quad (5)$$

The KR distribution

In this subsection, we give the new distribution called KR distribution. It is formulated using the Rani ($\theta > 0$) and the Kumaraswamy ($a = 1, b > 0$) distributions.

The KR density for $x > 0$ is expressed as

$$f(x) = b \left[\frac{\theta^5}{\theta^5 + 24} (\theta + x^4) e^{-\theta x} \right] \left[\left[1 + \frac{\theta x (\theta^3 x^3 + 4\theta^2 x^2 + 12\theta x + 24)}{\theta^5 + 24} \right] e^{-\theta x} \right]^{(b-1)} \quad (6)$$

The parameters vector of the KR distribution is $\Theta = (b, \theta)$. Moreover, using equation (3), equation (6) can be defined as:

$$f(x) = \left(\frac{\theta^5}{\theta^5 + 24} (\theta + x^4) e^{-\theta x} \right) \sum_{i=0}^{\infty} w_i \left(1 - \left[1 + \frac{\theta x (\theta^3 x^3 + 4\theta^2 x^2 + 12\theta x + 24)}{\theta^5 + 24} \right] e^{-\theta x} \right)^i, \quad (7)$$

where $w_i = (-1)^i b \binom{b-1}{i}$.

Remark 1: Let $X \sim KR(b, \theta)$. If $b=1$, then $X \sim R(\theta)$ (Figure 1).

We can find easily the cumulative distribution function (c.d.f) of the KR distribution

$$F(x) = 1 - \left[\left[1 + \frac{\theta x (\theta^3 x^3 + 4\theta^2 x^2 + 12\theta x + 24)}{\theta^5 + 24} \right] e^{-\theta x} \right]^b, \quad x > 0. \quad (8)$$

Using the c.d.f in order to obtain the Survival function (Figure 2):

$$S(x) = 1 - F(x) = \left[\left[1 + \frac{\theta x (\theta^3 x^3 + 4\theta^2 x^2 + 12\theta x + 24)}{\theta^5 + 24} \right] e^{-\theta x} \right]^b$$

and the Hazard function which is an important quantity characterizing life phenomena is given by

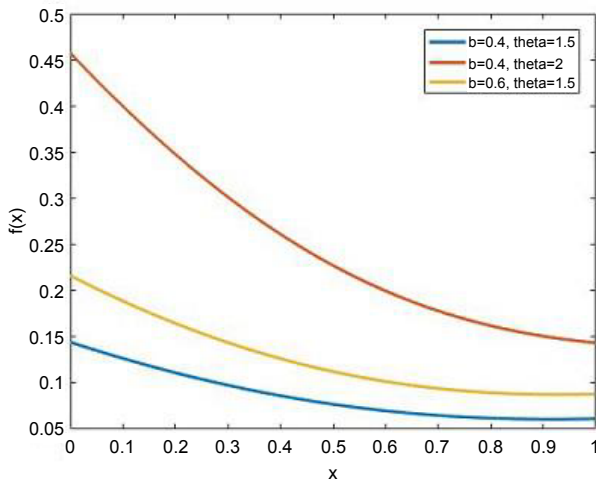


Figure 1: The KR densities of selected values of b and θ ...

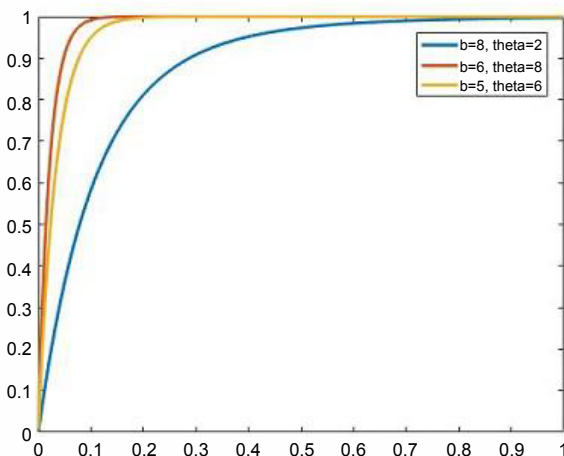


Figure 2: The cumulative distribution functions of selected values of b and θ ..

$$H(x) = \frac{f(x)}{1 - F(x)} = \frac{b \left[\frac{\theta^5}{\theta^5 + 24} (\theta + x^4) e^{-\theta x} \right]}{\left[1 + \frac{\theta x (\theta^3 x^3 + 4\theta^2 x^2 + 12\theta x + 24)}{\theta^5 + 24} \right] e^{-\theta x}}$$

Now, we shall introduce the following Theorem about the inverse cumulative KR distribution. The latter characterizes the proposed distribution.

Theorem 2.1: Let $Y \sim KR(b, \theta)$, then $x = F^{-1}(y)$ is the zero of the polynomial equation given by

$$\theta^4 x^4 + 4\theta^3 x^3 + Ax^2 + Bx + C, \quad (9)$$

where,

$$A = -\frac{\theta^7}{2} + (1 - (1 - \mu)^{\frac{1}{b}}) \left(\frac{\theta^7}{2} + 12\theta^2 \right),$$

$$B = (1 - (1 - \mu)^{\frac{1}{b}}) (\theta^6 + 24\theta) - \theta^6,$$

$$C = (24 + \theta^5) (1 - (1 - \mu)^{\frac{1}{b}})$$

and $\mu \in]0, 1[$.

Proof: In order to get the zero of equation (9), we shall introduce the following equation:

$$F(x) = 1 - (1 - G(x))^b = \mu, \quad \mu \in]0, 1[.$$

So,

$$1 - \left[1 + \frac{\theta x (\theta^3 x^3 + 4\theta^2 x^2 + 12\theta x + 24)}{\theta^5 + 24} \right] e^{-\theta x} = 1 - (1 - \mu)^{\frac{1}{b}}.$$

We have also,

$$(1 - \mu)^{\frac{1}{b}} - e^{-\theta x} \left[\frac{\theta^5 + 24 + \theta x (\theta^3 x^3 + 4\theta^2 x^2 + 12\theta x + 24)}{\theta^5 + 24} \right] = 0.$$

In the following, we get,

$$e^{\theta x} (1 - \mu)^{\frac{1}{b}} = \left[\frac{\theta^5 + 24 + \theta x (\theta^3 x^3 + 4\theta^2 x^2 + 12\theta x + 24)}{\theta^5 + 24} \right].$$

It is to be noted that the limited development to order 2 of exponential function at 0 can be written as:

$$e^{\theta x} = 1 + \theta x + \frac{(\theta x)^2}{2}.$$

Then, we obtain:

$$\left(1 + \theta x + \frac{(\theta x)^2}{2} \right) ((1 - \mu)^{\frac{1}{b}})^b = \frac{\theta^5 + 24 + \theta x (\theta^3 x^3 + 4\theta^2 x^2 + 12\theta x + 24)}{\theta^5 + 24}. \quad (10)$$

Applying some mathematical treatments on equation (10) yields,

$$\begin{aligned} & \theta^4 x^4 + 4\theta^3 x^3 + x^2 \left[-((1 - \mu)^{\frac{1}{b}}) \left(\frac{\theta^7}{2} + 12\theta^2 \right) \right] \\ & + \left[(1 - \mu)^{\frac{1}{b}} (\theta^6 + 24\theta) \right] x + (24 + \theta^5) (1 - (1 - \mu)^{\frac{1}{b}}) = 0. \end{aligned}$$

Proving that

$$\theta^4 x^4 + 4\theta^3 x^3 + Ax^2 + Bx + C = 0,$$

where,

$$A = -((1 - \mu)^{\frac{1}{b}}) \left(\frac{\theta^7}{2} + 12\theta^2 \right),$$

$$B = ((1 - \mu)^{\frac{1}{b}}) (\theta^6 + 24\theta)$$

and

$$C = (24 + \theta^5)(1 - (1 - \mu)^{\frac{1}{b}})$$

which completes this proof.

Parameters Estimations

In this section, the maximum likelihood method for the parameters estimation of the KR distribution is derived. For this reason, a brief description of the method for the fitting of the KR parameters is presented. The likelihood function for the parameters is given by

$$L = \prod_{i=1}^n \left(b \left[\frac{\theta^5}{\theta^5 + 24} (\theta + x_i^4) e^{-\theta x_i} \right] \left[\left[1 + \frac{\theta x_i (\theta^3 x_i^3 + 4\theta^2 x_i^2 + 12\theta x_i + 24)}{\theta^5 + 24} \right] e^{-\theta x_i} \right]^{(b-1)} \right). \quad (11)$$

The log-likelihood is formulated as:

$$\begin{aligned} \log(L) &= n \log(b) + \sum_{i=1}^n \log \left(\frac{\theta^5}{\theta^5 + 24} (\theta + x_i^4) e^{-\theta x_i} \right) \\ &+ \sum_{i=1}^n (b-1) \log \left(\left[1 + \frac{\theta x_i (\theta^3 x_i^3 + 4\theta^2 x_i^2 + 12\theta x_i + 24)}{\theta^5 + 24} \right] e^{-\theta x_i} \right). \end{aligned} \quad (12)$$

In order to get the maximum likelihood estimators of (b, θ) , the first order partial derivatives function of $\log(L)$ is calculated, with respect to the KR parameters which are given respectively by:

$$\frac{\partial \log(L)}{\partial b} = \frac{n}{b} + \sum_{i=1}^n \log \left(\left[1 + \frac{\theta x_i (\theta^3 x_i^3 + 4\theta^2 x_i^2 + 12\theta x_i + 24)}{\theta^5 + 24} \right] e^{-\theta x_i} \right). \quad (13)$$

$$\begin{aligned} \frac{\partial \log(L)}{\partial \theta} &= \frac{5n}{\theta} + \sum_{i=1}^n \frac{1}{\theta + x_i^4} - b \sum_{i=1}^n x_i - nb \frac{5\theta^4}{\theta^5 + 24} \\ &+ (b-1) \sum_{i=1}^n \left[\frac{5\theta^4 + 4\theta^3 x_i^4 + 12\theta^2 x_i^3 + 24\theta x_i^2 + 24x_i}{\theta^5 + 24 + \theta^4 x_i^4 + 4\theta^3 x_i^3 + 12\theta^2 x_i^2 + 24\theta x_i} \right]. \end{aligned} \quad (14)$$

By setting these above equations to zero, we get

$$b = \frac{-n}{\sum_{i=1}^n \log \left(\left[1 + \frac{\theta x_i (\theta^3 x_i^3 + 4\theta^2 x_i^2 + 12\theta x_i + 24)}{\theta^5 + 24} \right] e^{-\theta x_i} \right)} \quad (15)$$

and

$$\theta = \frac{5n}{-\sum_{i=1}^n \left(\frac{1}{\theta + x_i^4} \right) + b \sum_{i=1}^n x_i + nb \left(\frac{5\theta^4}{\theta^5 + 24} \right) - (b-1) \sum_{i=1}^n \left[\frac{5\theta^4 + 4\theta^3 x_i^4 + 12\theta^2 x_i^3 + 24\theta x_i^2 + 24x_i}{\theta^5 + 24 + \theta^4 x_i^4 + 4\theta^3 x_i^3 + 12\theta^2 x_i^2 + 24\theta x_i} \right]}$$

Hence, the maximum likelihood estimators of b and θ can be obtained by using the fixed point method or Newton — Raphson method or other numerical methods.

Simulation Studies

Let us consider some simulation studies to illustrate the usefulness of the new distribution. By applying the inverse transformation technique, we simulated two random samples having size 5000 from KR distribution, with the parameter vectors $(b=2, \theta=3)$ and $(b=5, \theta=2)$ respectively (Figures 3 and 4).

From the above Histograms and compared to the curve of the proposed density, it is worth noting that we have generated samples from a KR distribution.

Application

Our basic objective in this section is to motivate the use of the KR distribution by showing a successful application to modeling lifetime data set given [9] which represents the relief times of twenty patients receiving an analgesic. The data are as follows: 1.1,1.4,1.3,1.7,1.9,1.8,1.6,2.2,1.7,2.7,4.1,1.8,1.5,1.2,1.4,3.0,1.7,2.3,1.3,2.0

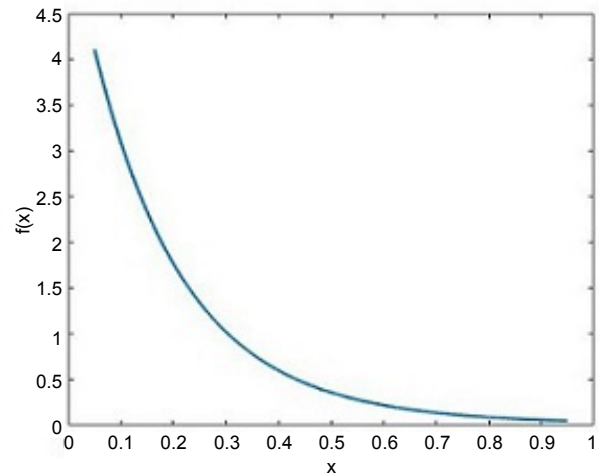


Figure 3: Histogram of random samples from a KR distribution and a curve of the density: $(b = 2, \theta = 3)$.

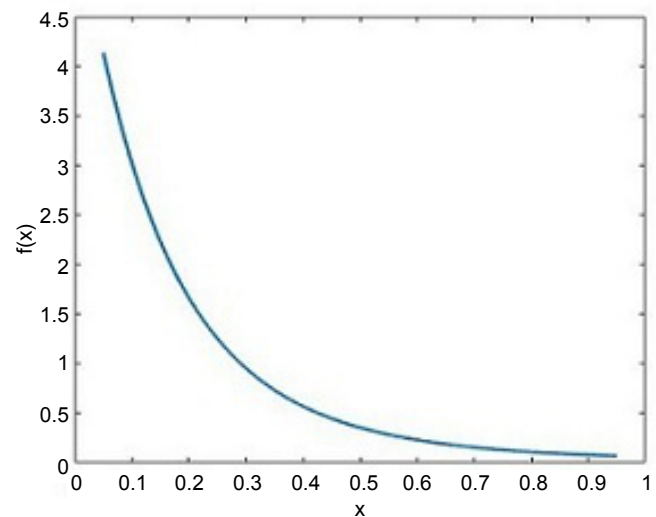


Figure 4: Histogram of random samples from a KR distribution and a curve of the density: $(b = 5, \theta = 2)$.

We use the lifetime data set given above to evaluate the effectiveness of the new distribution with three distributions: Rani, Akash and Sujatha distributions [7,10,11]. To assess the goodness-of-fit of above models, we used the Akaike information criterion (AIC), the Akaike Information Criterion Corrected (AICC), the Bayesian information criterion (BIC) and the Log-likelihood L , which are calculated according to the following formula:

$$AIC = -2\log(L) + 2k,$$

$$AICC = AIC + \frac{2k(k+1)}{n-k-1}$$

and

$$BIC = -2\log(L) + k\log(n)$$

where, n is the number of observations in the sample and k represents the number of parameters.

The maximum likelihood estimates of the parameters $(\hat{b}, \hat{\theta})$, the Akaike Information Criterion (AIC) the Akaike Information Criterion Corrected (AICC), the Bayesian Information Criterion (BIC) and the log-likelihood L values for the corresponding distributions are reported in Table 1.

According to the small values of both AIC , BIC and $AICC$ the proposed distribution performs better than the other candidate models. In addition, referring to the big L values for the KR distribution, we infer that the latter

Table 1: The AIC, AICC, BIC and L values for the different distributions.

Distributions	Maximum Likelihood estimates	AIC	AICC	BIC	L
KR	$\hat{b} = 3.218, \hat{\theta} = 0.2501$	58.2524	58.9582	60.2439	-27.1262
Rani	$\hat{\theta} = 1.4727$	65.4474	65.6697	66.4432	-31.7237
Akash	$\hat{\theta} = 1.7608$	67.4118	67.6340	68.4075	-32.7059
Sujatha	$\hat{\theta} = 1.6$	66.8815	66.8824	73.2935	-32.4407

provides a good fit for the given data and it proves therefore to be the more appropriate model.

Conclusion

In this paper, we set forerend a new distribution namely the KR distribution. It is formulated using the Kumerswamy-generalized distribution and the Rani distribution. We provide a mathematical treatment of the new distribution including cumulative, Hazar and Survival functions. We discuss a maximum likelihood estimation of the model's parameters. An application of the new distribution to a real data set is presented to demonstrate that it can be used quite effectively to provide better fits than other available models. We aspire this generalization may attract wider applications in survival analysis.

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